On the greatest zero of an orthogonal polynomial. I

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Dedicated to Prof. Béla Szőkefalvi-Nagy on the occasion of his 60th birthday

0. Introduction

Let w(x) $(-\infty < x < \infty)$ be an *even* weight function, and let $\{p_n(w; x) = \gamma_n(w)x^n + \cdots; n=0, 1, \ldots\}$ be the sequence of orthonormal polynomials with respect to w, i.e.

(1)
$$\int_{-\infty}^{\infty} p_m(w; x) p_n(w; x) w(x) dx = \begin{cases} 0 & (m \neq n), \\ 1 & (m = n). \end{cases}$$

Moreover, let $X_n(w) = x_{1n}(w)$ be the greatest zero of $p_n(w; x)$. In part 1 of the present note we express the order of $X_n(w)$ with the aid of the sequence $\{\gamma_v(w)\}$ (see Theorem 1). After deducing some lemmas in part 2, we apply this result in part 3 to the weight

(2)
$$w_{\varrho,2k}(x) = |x|^{\varrho} e^{-x}$$

where $\varrho \ge 0$ and k is a positive integer. We prove the estimate $\gamma_{\nu-1}(w_{\varrho,2k})/\gamma_{\nu}(w_{\varrho,2k}) = O(\nu^{1/2k})$ which seems far from being trivial and conclude from it that

(3)
$$X_n(w_{\varrho,2k}) \sim n^{1/2k} \sim \gamma_{n-1}(w_{\varrho,2k})/\gamma_n(w_{\varrho,2k}).$$

The relation (3) has several interesting implications in approximation theory; we hope to return to them soon.

1. An inequality on $X_n(w)$

Theorem 1. For every even weight function w(x) we have

(4)
$$\max_{1 \le k \le n-1} \frac{\gamma_{k-1}(w)}{\gamma_k(w)} \le X_n(w) \le 2 \max_{1 \le k \le n-1} \frac{\gamma_{k-1}(w)}{\gamma_k(w)}.$$

Remarks. a) Let $w_0(x) = (1-x^2)^{-1/2}$ with support [-1, 1]. Then the first three orthogonal polynomials are $\frac{1}{\sqrt{\pi}}, \sqrt{\frac{2}{\pi}}x, \sqrt{\frac{2}{\pi}}(2x^2-1)$, i.e. $\gamma_0(w_0)/\gamma_1(w_0) =$

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 $=\frac{1}{\sqrt{2}}=X_2(w_0)$. This example shows that the left-hand part of inequality. (4) is precise.

b) In case $w_{02}(x) = e^{-x^2}$ the orthogonal polynomials $p_n(w)$ are the orthonormal Hermite polynomials $h_n(x)$, so that $\gamma_{n-1}(w_{02})/\gamma_n(w_{02}) = \sqrt{\frac{n}{2}}$ and $X_n(w_{02}) \approx \sqrt{2n}$. This example shows that the factor 2 on the right-hand side of (4) can not be replaced by any smaller constant.

Proof. By a classical result of P. L. CHEBYCHEV (see G. SZEGŐ [2], 7.7.2) we have

(5)
$$X_{n}(w) = \max \frac{\int_{-\infty}^{\infty} x [P_{n-1}(x)]^{2} w(x) dx}{\int_{-\infty}^{\infty} [P_{n-1}(x)]^{2} w(x) dx},$$

where $P_{n-1}(x)$ runs over all polynomials of degree $\leq n-1$. Let us represent $P_{n-1}(x)$ as

(6)
$$P_{n-1}(x) = \sum_{j=0}^{n-1} c_j p_j(w; x).$$

We recall that by the recursion formula applied to even w we have

(7)
$$xp_{j}(w;x) = \frac{\gamma_{j}(w)}{\gamma_{j+1}(w)}p_{j+1}(w;x) + \frac{\gamma_{j-1}(w)}{\gamma_{j}(w)}p_{j-1}(w;x)$$

(see e.g. G. FREUD [1], § I. 2).

Inserting (6) into (5) and taking (1) and (7) into consideration we obtain

(8)
$$X_n(w) = 2 \max \frac{\sum_{k=1}^{n-1} \frac{\gamma_{k-1}(w)}{\gamma_k(w)} c_{k-1} c_k}{\sum_{k=0}^{n-1} c_k^2},$$

where all the c_k (k = 0, 1, ..., n-1) run idependently over the reals. Inserting $c_{i-1} = c_i = 1$ and $c_k = 0$ if $k \neq j-1$, j into the expression on the right of (8), we obtain

(9)
$$X_n(w) \ge \frac{\gamma_{j-1}(w)}{\gamma_j(w)}$$
 $(j = 1, 2, ..., n-1).$

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In turn, by Cauchy's inequality for every $\{c_k\}$ we have

$$\sum_{k=1}^{n-1} \frac{\gamma_{k-1}(w)}{\gamma_k(w)} c_{k-1} c_k \leq \max_{1 \leq k \leq n-1} \frac{\gamma_{k-1}(w)}{\gamma_k(w)} \sum_{k=0}^{n-1} |c_{k-1} c_k| \leq \sum_{1 \leq k \leq n-1} \frac{\gamma_{k-1}(w)}{\gamma_k(w)} \sum_{k=1}^{n-1} c_k^2.$$

The left-hand side of (4) is implied by (9) and the right hand side of (4) is a consequence of (8) and (10), and so Theorem 1 is proved.

2. Lemmata

Let

(11)
$$w_{\varrho\beta}(x) = |x|^{\varrho} e^{-|x|^{\beta}} \qquad (-\infty < x < \infty)$$

Lemma 1. For every $\varrho \ge 0$ and $\beta > 0$ we have

(12)
$$\frac{\gamma_n(w_{\varrho\beta})}{\gamma_{n-1}(w_{\varrho\beta})} = \frac{\beta}{n+\Delta_n \varrho} \int_{-\infty}^{\infty} p_n(w_{\varrho\beta}; x) p_{n-1}(w_{\varrho\beta}; x) x^{-1} |x|^{\beta} w_{\varrho\beta}(x) dx$$

where

(13)
$$\Delta_n = \frac{1}{2} [1 + (-1)^{n+1}].$$

Proof. We have

$$\int_{-\infty}^{\infty} p'_{n}(w_{\varrho\beta}; x) p_{n-1}(w_{\varrho\beta}; x) w_{\varrho\beta}(x) dx =$$

$$= \int_{-\infty}^{\infty} [n\gamma_{n}(w_{\varrho\beta}) x^{n-1} + \cdots] p_{n-1}(w_{\varrho\beta}; x) w_{\varrho\beta}(x) dx =$$

$$= \int_{-\infty}^{\infty} \left[n \frac{\gamma_{n}(w_{\varrho\beta})}{\gamma_{n-1}(w_{\varrho\beta})} p_{n-1}(w_{\varrho\beta}; x) + P_{n-2}(x) \right] p_{n-1}(w_{\varrho\beta}; x) w_{\varrho\beta}(x) dx,$$

where P_{n-2} is a polynomial of degree $\leq n-2$. Applying the orthogonality relations (1), we get

(14)
$$n\frac{\gamma_n(w_{\varrho\beta})}{\gamma_{n-1}(w_{\varrho\beta})} = \int_{-\infty}^{\infty} p'_n(w_{\varrho\beta}; x) p_{n-1}(w_{\varrho\beta}; x) w_{\varrho\beta}(x) dx.$$

Partial integration gives

$$\int_{-\infty}^{\infty} p'_{n}(w_{e\beta}; x) p_{n-1}(w_{e\beta}; x) w_{e\beta}(x) dx = -\int_{-\infty}^{\infty} p_{n}(w_{e\beta}; x) [p_{n-1}(w_{e\beta}; x) w_{e\beta}(x)]' dx =$$

(15)
$$= \beta \int_{-\infty}^{\infty} p_n(w_{\varrho\beta}; x) p_{n-1}(w_{\varrho\beta}; x) x^{-1} |x|^{\beta} w_{\varrho\beta}(x) dx - \rho \int_{-\infty}^{\infty} p_n(w_{\varrho\beta}; x) p_{n-1}(w_{\varrho\beta}; x) x^{-1} w_{\varrho\beta}(x) dx,$$

since $\int p_n(w)p'_n(w)w dx = 0$ by (1).

If n is even, $p_{n-1}(w_{\rho\beta}; x)$ is odd, and so $x^{-1}p_{n-1}(w_{\rho\beta}; x)$ is a polynomial of degree n-2. Consequently, the second integral on the right of (15) vanishes by (1). In this way, from (15) we obtained

6)
$$\int_{-\infty}^{\infty} p'_{n}(w_{e\beta}; x) p_{n-1}(w_{e\beta}; x) w_{e\beta}(x) dx =$$
$$= \beta \int_{-\infty}^{\infty} p_{n}(w_{e\beta}; x) p_{n-1}(w_{e\beta}; x) x^{-1} |x|^{\beta} w_{e\beta}(x) dx. \quad (n \text{ is even})$$

Let now *n* be odd. Then $p_n(w_{o\beta}; x)$ is odd, and so

$$x^{-1}p_{n}(w_{\varrho\beta};x) = \frac{\gamma_{n}(w_{\varrho\beta})}{\gamma_{n-1}(w_{\varrho\beta})}p_{n-1}(w_{\varrho\beta};x) + P_{n-2}(x),$$

where $P_{n-2}(x)$ is a polynomial of degree $\leq n-2$. Using the orthogonality relation (1) we see that

(17)
$$\int_{-\infty}^{\infty} p_n(w_{\varrho\beta}; x) p_{n-1}(w_{\varrho\beta}; x) x^{-1} w_{\varrho\beta}(x) dx = \frac{\gamma_n(w_{\varrho\beta})}{\gamma_{n-1}(w_{\varrho\beta})}. \quad (n \text{ is odd})$$

From (14), (15), (16), and (17) we see that (12) holds for both even and odd integers n. Q.E.D.

Lemma 2. For every positive integer k we have

(18)
$$\left[\frac{\gamma_{n-1}(w_{\varrho,2k})}{\gamma_n(w_{\varrho,2k})}\right]^{2k} \leq \frac{n+\varrho\Delta_n}{2k}.$$

Remark. For k=1, $\varrho=0$ we have equality in (18).

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Proof. We infer by induction from the recursion formula (7) that, for every positive integer l and every *even* w, we have

(19)
$$x^{l} p_{n}(w; x) = \sum_{j=0}^{n+l} A_{n,l,j}(w) p_{j}(w; x),$$

where all coefficients $A_{n,l,j}(w)$ are nonnegative. By (7) we have

(20)
$$A_{n,1,n-1}(w) = \frac{\gamma_{n-1}(w)}{\gamma_n(w)}.$$

Moreover, by a repeated application of the recursion formula (7) we obtain (21)

$$A_{n-1,2,n-1}(w) = \int_{-\infty}^{\infty} x^2 p_{n-1}^2(w;x) w(x) dx = \frac{\gamma_{n-1}^2(w)}{\gamma_n^2(w)} + \frac{\gamma_{n-2}^2(w)}{\gamma_{n-1}^2(w)} \ge \frac{\gamma_{n-1}^2(w)}{\gamma_n^2(w)}$$

Multiplying (19) by x^2 and then applying the special case l=2 of the same formula to the right-hand side, we get

(22)
$$A_{n,l+2,n-1}(w) \ge A_{n,l,n-1}(w) A_{n-1,2,n-1}(w).$$

From (21) and (22) we infer by induction that

$$\int_{-\infty}^{\infty} x^{2s-1} p_n(w; x) p_{n-1}(w; x) w(x) dx = A_{n, 2s-1, n-1}(w) \ge \int_{-\infty}^{\infty} x^{2s-1} (w) e^{-1} dx$$

(23)

$$\geq \left[\frac{\gamma_{n-1}(w)}{\gamma_n(w)}\right]^{2s-1} \qquad (s=1,2,\ldots).$$

Let us now insert $\beta = 2k$ in (12) and $w = w_{\varrho, 2k}$ in (23). Combining the two formulas so obtained we get (18). Q.E.D.

We introduce the moments

(24)
$$\mu_r(w) = \int_{-\infty}^{\infty} x^r w(x) dx \qquad (r = 0, 1, ...).$$

Lemma 3. For every even w, we have

(25)
$$[X_n(w)]^2 \ge \mu_{2n-2}(w)/\mu_{2n-4}(w).$$

Proof. Denoting by $X_n(w) = x_{1n} > x_{2n} > \cdots > x_{nn} = -X_n(w)$ the zeros of $p_n(w; x)$, by the Gauss—Jacobi quadrature formula we have

$$\mu_{2n-2}(w) = \sum_{j=1}^{n} \lambda_{jn}(w) x_{jn}^{2n-2} \leq [X_n(w)]^2 \sum_{j=1}^{n} \lambda_{jn}(w) x_{jn}^{2n-4} = [X_n(w)]^2 \mu_{2n-4}(w).$$

Q.E.D.

We can also see that the sign of equality is valid in (25) iff n=2.

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Example 13. Estimates for $X_n(w_{e,2k})$

Theorem 2. For every $\varrho \ge 0$ and $\beta > 0$ we have

(26)
$$\lim_{k \to \infty} n^{-1/\beta} X_n(w_{\varrho\beta}) \ge (2/\beta)^{1/\beta}$$

Proof. We have

(27)
$$\mu_{2r}(w_{\varrho\beta}) = 2\int_{0}^{\infty} x^{2r+\varrho} e^{-x\beta} dx = \Gamma\left(\frac{2r+\varrho+1}{\beta}\right) \qquad (r=0,1,\ldots).$$

Insert (27) in (25) and apply Stirling's formula to get the desired result.

Theorem 3. For every $\varrho \ge 0$ and every positive integer k we have

(28)
$$X_n(w_{\varrho, 2k}) \le 2(n/2k)^{1/2k}$$

Remark. We have $X_n(w_{0,2}) \approx \sqrt{2n}$. So the factor on the right of (28) cannot be replaced by any constant smaller than 2.

Proof. This is a consequence of Theorem 1 and Lemma 2. We see from Theorem 2 and Theorem 3 that

(29) $X_n(w_{q,2k}) \sim n^{1/2k}$

holds for every $\varrho \ge 0$ and every positive integer k.

Theorem 4. For every $\varrho \ge 0$ and every positive integer k we have

(30)
$$\frac{\gamma_{n-1}(w_{e,2k})}{\gamma_n(w_{e,2k})} \ge 2^{-2k+1} \left(\frac{n}{2k}\right)^{1/2k} \left(1+\frac{k}{n}\right)^{\frac{2k-1}{2k}}.$$

Remark. From (30) and the combination of (28) and the left hand side of (4) we see that

(31)
$$\frac{\gamma_{n-1}(w_{\varrho,2k})}{\gamma_n(w_{\varrho,2k})} \sim n^{1/2k}.$$

Proof. Consider formula (12). The expression $p_n(w_{\varrho,2k}; x)p_{n-1}(w_{\varrho,2k}; x)x^{2k-1}$ is a polynomial of degree 2n+2k-2 < 2(n+k)-1. Consequently the integral in (12) can be calculated by the Gauss—Jacobi quadrature formula over the zeros of $p_{n+k}(w_n; x)$:

$$\frac{n}{2k} \frac{\gamma_{n}(w_{\varrho,2k})}{\gamma_{n-1}(w_{\varrho,2k})} \leq \frac{n+\varrho\Delta_{n}}{2k} \frac{\gamma_{n}(w_{\varrho,2k})}{\gamma_{n-1}(w_{\varrho,2k})} =$$

$$= \sum_{j=1}^{n+k} \lambda_{j,n+k}(w_{\varrho,2k}) x_{j,n+k}^{2k-1} p_{n}(w_{\varrho,2k}; x_{j,n+k}) p_{n-1}(w_{\varrho,2k}; x_{j,n+k}) \leq$$

$$\leq [X_{n+k}(w_{\varrho,2k})]^{2k-1} \left\{ \sum_{j=1}^{n+k} \lambda_{j,n+k}(w_{\varrho,2k}) p_{n}^{2}(w_{\varrho,2k}; x_{j,n+k}) \times \right\}^{1/2} = [X_{n+k}(w_{\varrho,2k})]^{2k-1},$$

since by the quadrature formula we have

$$\sum_{j=1}^{n+k} \lambda_{j,n+k}(w_{\varrho,2k}) p_r^2(w_{\varrho,2k}; x_{j,n+k}) = \int_{-\infty}^{\infty} p_r^2(w_{\varrho,2k}; x) w_{\varrho,2k}(x) dx = 1 \quad (r = n-1, n).$$

Inserting estimate (28) into the right-hand side of (32), we obtain the desired estimate (30) after reshuffling the factors. Q.E.D.

Literature

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